## Final Report: Sampling for the European Social Survey

## 1. Introduction

The European Social Survey is a new, academically-driven social survey designed to chart and explain the attitudes, beliefs and behaviour patterns of Europe's diverse populations. The survey covers 23 nations and aims at asserting the most rigorous methodologies. Therefore, a methodological overhead with ten Work packages has been installed. ${ }^{1}$ One of them deals with sampling: A panel of experts was signing off all sampling designs in advance of the fielding periods in the different countries ${ }^{2}$.

In this report we firstly want to explain the principles of sampling for the ESS. Secondly we want to show which way we organised the process of "Signing off" the sampling designs. Thereafter, a short description of the design of each participating country is given. Finally, we present some comprising ideas.

## 2. Principles and requirements of sampling for the ESS

The objective of the Sampling Work package was the "design and implementation of workable and equivalent sampling strategies in all participating countries". This concept stands for random samples with comparable estimates. From the statistical point of view full coverage of the population, low non-response rates and consideration of design effects are prerequisites for the comparability of unbiased or at least minimum biased estimates. In the following we want briefly to describe these requirements and to show some examples, how the requirements could be kept in the practices of the individual countries.

### 2.1 Basic principles for sampling in cross-cultural surveys

Kish (1994, p. 173) provides our starting point: "Sample designs may be chosen flexibly and there is no need for similarity of sample designs. Flexibility of choice is particularly advisable for multinational comparisons, because the sampling resources differ greatly between countries. All this flexibility assumes probability selection methods: known probabilities of selection for all population elements." Following this, an optimal sampling design for crosscultural surveys should consist of the best random practice used in each participating country. The choice of a specific design depends on the available frames, experiences and, of course,

[^0]also the costs in the different countries (Häder/Gabler 2003). If adequate estimators are chosen the resulting values can be compared. This comparability stands for equivalence and is the goal of the sampling strategy and its implementation for the ESS.

## 2. 2 Discussion of standards set in the Specification for participating countries ${ }^{3}$

Only random samples provide a theoretical basis, which allows us to infer from the sample to the population or sub-sets of this. As design-based inference is one important goal in the project, probability samples are required. However, this is related to other requirements:

- full coverage of the target population
- high response rates
- no substitution
- the same minimum effective sample sizes in participating countries (ESS: $\mathrm{n}_{\text {eff }}=1,500$ or 800 where population is smaller than 2 million inhabitants) and a minimum net sample size of $n_{\text {net }}=2,000$.

These requirements can only be sensibly discussed in the context of random samples. They form a theoretical system that in the end ensures equivalence.

## Full coverage of the residential population

An important step in planning a survey is the definition of the population under study (target population). In the case of the ESS it contains in each country persons 15 years or older who are resident within private households, regardless of nationality and citizenship, language ${ }^{4}$ or legal status. This definition applies to all participating countries. Thus, every person with the defined characteristics should have a non zero chance of being selected. This implies, that the more completely the frame covers the persons belonging to the target population, the better the resulting sample will be. However, the quality of the frames - e.g. coverage, updating and access - differs from country to country. Therefore, frames have to be evaluated carefully. The results of these evaluations are documented and have to be taken into account when the data are analysed.
Among others, we found the following kinds of frames:
a) countries with reliable lists of residents that are available for social research such as the Danish Central Person Register that has approximately $99.9 \%$ coverage of persons resident in Denmark

[^1]b) countries with reliable lists of households that are available for social research such as the "SIPO" database in the Czech Republic, that is estimated to cover $98 \%$ of households
c) countries with reliable lists of addresses that are available for social research such as the postal delivery points from "PTT-afgiftenpuntenbestand" in the Netherlands
d) countries without reliable and/or available lists such as Portugal or France

In all cases, fortunately, there is some aggregated demographic information available that can be used for the sampling strategy. The update of this information varies to some extent, from some months until a few years.

Drawing a sample is more complicated if no registers (lists) are available (group d). In this instance area based designs are usually applied, in which the selection of municipalities forms the first stage and the selection of households within these municipalities the second stage. Because no sampling frames are available, the crucial problem is the selection of households. There are two main ways to go about this. The first is to list all the addresses within certain areas of each selected community. The target households are drawn from these lists. It is possible to assess this procedure as one way of drawing a random sample, even if one which is fairly strongly clustered. A design of this kind is applied in Greece. Another frequently used way to find target households is the application of random route elements. The question here, however, is the extent to which random routes can be judged to be "strictly random". That depends on both, the definition of the rules for the random walk and the control of the interviewers by the fieldwork organisation in order to minimise the interviewer's influence on the selection of respondents. In Austria, e.g., there is a design with a random route element. The survey institute together with the National Co-ordinator of the ESS operationalised the general rules for various household types (like large apartment buildings, small houses within densely populated areas, houses in the countryside, etc). Moreover, all selected households will be checked by the supervising team. This approach was convincing for the sampling expert panel.

Even in countries where reliable frames exist, some problems had to be solved. For example, in Italy there is an electoral register available. But it contains, of course, only persons 18 years or older. Therefore, it had to be used as a frame of addresses. In Ireland, we found the same situation.

People with illegal status will be underrepresented because they are not registered. Such systematic losses because of undercoverage cannot be ruled out in practice. However, they must be documented carefully.

## Response rates

Non-response is the next problem for the representativeness of the target population in the sample. A carefully drawn gross sample from a perfect frame can be worthless if non-contacts and refusals lead to systematic biases. Therefore, it is of essential importance to plan and implement a sufficient number of contacts as well as appropriate fieldwork strategies for the persuasion of the target persons to participate in the survey. For the ESS a target response rate of $70 \%$ has been fixed. This may be particularly challenging for some countries where response rates between 40 and 55 percent are common (Lyberg 2000). Nevertheless, all efforts should be done to avoid non-response because it includes the danger of biased samples, and cell weighting is not as global a means of "repairing" samples as is sometimes argued. The expected response rates among the ESS countries range from 30\% (Luxembourg) to $75 \%$ (e.g. Sweden and Denmark). In Switzerland (40\%) a special methodological experiment is integrated to study possibilities of achieving a higher response rate. Most countries hope to get a response rate of about $70 \%$. How realistic this hope is will be seen after the end of the fielding process. In any case, many different techniques for increasing the response rates such as advance letters, toll-free telephone numbers for potential respondents to contact, extra training of interviewers in response-maximisation techniques and doorstep interactions are applied in all countries.

## Substitution

Connected with non-response and the resulting fear of biased samples is the problem of substituting non-cooperative or not reachable primary sampling units, households or target persons by others. This practice is sometimes applied to get "better samples". However, substitution cannot be defended on theoretical grounds, because co-operative and easy-toreach people would be over-represented in the sample. Furthermore, uncertainties concerning the inclusion probabilities are the result. Another important disadvantage of substitution in the field is that it tends to reduce the extent of interviewer efforts to gain a response at the original addresses/households (Elliot 1993). Thus, for the ESS substitution of non-responding households or individuals (whether 'refusals' or 'non-contacts') is not permitted. One exception is as follows: at the first stage of the sampling process administrative considerations may mean that addresses cannot be obtained for specific areas, which originally belonged to the sample. In these exceptional cases it is allowed to replace the areas with areas of the same strata.

Design effects indicate the precision of estimates (Kish 1987). In this sense they are criteria for the quality of the estimates and have influence on their equivalence. We chose a model based approach for the estimation of the design effects (Gabler/Häder/Lahiri 1999) for determining the sample sizes for each country. Depending on the available frames and the available funding ${ }^{5}$ we have more or less complex sampling designs, ranging from simple random sampling (Finland) to multistage stratified and clustered sampling (e.g. Poland, Spain). These different designs lead to various design effects. We found the following:

- Design effect due to differing selection probabilities (DEFFp)

In some countries, it is necessary to select the sample in stages, with the penultimate stage being residential households. In this case, each person's selection probability depends on the respective household size. Design effects of this kind have to be predicted e.g. for Ireland, Israel, Portugal, or the Czech Republic. Another reason why differing selection probabilities are used is that minority groups are over-sampled. Examples for this are Germany, where the East German population is over-sampled or Israel, where the Arab population is oversampled. A third design effect due to differing selection probabilities occurs in countries with the above described multi-stage designs where the PSUs are selected proportional to the size of individuals but not to the size of households such as in Israel or Ireland.

To give an idea of the sizes of the predicted design effects due to unequal selection probabilities here are some values: 1.33 for Ireland, 1.25 for Switzerland, 1.19 for the Netherlands, 1.16 for Czech Republic and 1.01 for Italy.

- Design effect due to clustering (DEFFc)

The cluster size of the selection units and the intra-class correlation also influence the design effect. The cluster size should be chosen as small as possible since: The larger the average cluster sizes are, the lower the effective sample size is and the more interviews have to be conducted to reach the minimum effective sample size of 1,500 . In that sense a large number of selection units with only a few interviews in each is the goal - but cannot be reached in each country because of funding restrictions.
In some countries calculations were made to estimate intra-class correlation coefficients from earlier surveys with similar variables. If there was no available empirical evidence at all upon which to base an estimate of the intra class correlation coefficient, then a value of 0.02 has
been suggested in the instruction material. This value was used in most countries. Because of the varying cluster sizes and intra-class correlations we estimated fairly different DEFFc. Here are some examples: 1.38 for Germany, 1.22 for Israel, 1.20 for Ireland, 1.18 for Switzerland, 1.15 for Greece and 1.14 for Poland.

The total design effect is the product of the design effect due to differing selection probabilities (DEFFp) and the design effect due to clustering (DEFFc). These total design effects vary largely between the countries and the net sample sizes had to exceed the target of 2,000: For instance, in Ireland will be conducted 2,400 interviews (total design effect: 1.6), in Switzerland there will be 2,205 interviews (total design effect: 1.47) and in Spain there is a net sample size of 2,085 (total design effect: 1.39).

### 2.3 Summary

Comparability of sampling means that the national surveys must provide estimates that are subject to minimal bias of parameters of the equivalent populations. The basic requirement to use probability samples together with the additional requirements discussed in this paper leads theoretically to comparable estimates with the same accuracy level for each country. However, in the end the quality of the data depends also on the implementation process, e.g. the practical applications ${ }^{6}$. Therefore, this process has to be monitored carefully.
For the ESS we could find ways to develop probability sampling methods in all participating countries that satisfy the specified rules.

## 3. The process of "Signing off" the sampling designs

In the following we want to describe how we organised the "signing off" of the sampling designs in the different countries. It may be of interest because an approach like that has never been used before in planning a cross-national survey.
Firstly, an expert panel with five members was installed ${ }^{7}$.

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Each of the experts was assigned about five countries to liase and support. The panellists contacted "their" National Co-ordinators asking for information about the foreseen sampling design. Following that a process of co-operation between the National Co-ordinators, the survey organisations and the sampling experts started. In many countries completely new designs had to be developed to meet the strict methodological requirements of the ESS. In other countries, it was only a matter of clarifying details. In particular, support in calculating the effective sample sizes often was necessary. The sampling experts also had to visit some countries for a detailed discussion of problems such as an expected low response rate, the selection of a well suited survey organisation or the development of a completely new design. After the clarification of all questions the design of a country was "Ready for Signing off". The Expert panel had developed a form, where details of the design of each country had to be filled in. This task was done by the experts for "his/her" countries. For that the expert used the information from the previous discussion with the National Co-ordinator and the Survey Organisation in the country concerned. This procedure ensured that the terms used in the forms are standardised "statistical language" and that the design was really clear defined. Then the expert presented the form to the other panellists. If all of them agreed the design was "signed off". Otherwise, the discussion with the National Co-ordinator had to carry on. Thus, the decision of signing off a design was always made by the whole team together.
At the third meeting of the expert panel on $15^{\text {th }}$ October 2002 in Helsinki most countries could be signed off. However, a few countries without experience in probability sampling kept the expert panel occupied till May 2003 when we could sign off the design of the last country. At the meeting on $19^{\text {th }}$ May in London the National Co-ordinators evaluated the work of the sampling expert panel as very helpful.

## Time schedule:

- $17^{\text {th }}$ December 2001 in London: First sampling panel meeting. Discussion of the general approach for the process of "signing off" the sampling designs of the different countries
- $18^{\text {th }}$ December 2001 in London: Kick-off meeting of the National Co-ordinators. Explanation of the principal requirements of sampling for the ESS, introduction of the expert panel.
- $6^{\text {th }}$ and $7^{\text {th }}$ May 2002 in Mannheim: Second sampling panel meeting. Discussion of current status of ESS countries and outstanding issues.
- $26^{\text {th }}$ August 2002 in Copenhagen: Presentation of a paper on sampling for the ESS at the International Conference on Improving Surveys.
- $15^{\text {th }}$ October 2002 in Helsinki: Third expert panel meeting. Signing off most countries.
- May 2003: 'Signing off' of the $23^{\text {rd }}$ country


## 4. Description of the designs and deviations from the "Specifications for participating countries"

In the following section we describe the designs of the different countries in a standardized form. Special emphasis is placed on the nature of the frames, the sampling procedure and the computation of the sample sizes. For deviations from the "Specifications for participating countries" reasons are given. The forms are in alphabetic order.
A summary of the most important figures such as predicted design effects, anticipated response rates and number of Primary Sampling Units is given in Appendix 1.

Country:
NC :
Survey Institute:
Expert:

## Austria

Karl H. Müller
Institute for Panel Research (IPR), Richard Költringer
Sabine Häder

| Target Population, <br> Population coverage | Persons 15 years or older who are resident within private households in Austria, <br> regardless of nationality and citizenship, language or legal status. Homeless and <br> institutional populations are excluded. |
| :--- | :--- |
| Problems | People not understanding German language are excluded. |


| Sampling frame | For Stratification: Census 2001 <br> For selection of households: Austrian Telephone Book (CD), which is updated four <br> times a year |
| :--- | :--- |
| Problems | The Austrian Telephone Book covers only about $90 \%$ of the households. Not covered <br> are households without any telephone and households with secret numbers. |


| Sampling design | Stratified three stage probability sampling <br> Regional stratification: 363 strata in 121 districts $* 3$ classes of population sizes of <br> municipalities (small: $<2,500$, medium: $<10,000$, large: $\geq 10,000$ ) |
| :--- | :--- |
| Stage 1: Primary sampling units: 324 clusters in 251 municipalities (incl. 23 districts of <br> Vienna). The number of clusters in a stratum is proportional to the size of its <br> population (15 years and older). The allocation is done by controlled rounding (Cox <br> 1987). The selection within a stratum is done by systematic proportional-to-size <br> random sampling. <br> Stage 2: In each of the clusters 12 individuals are selected for the gross sample. For |  |
| that, in each cluster 6 addresses of households are drawn from the telephone book. |  |
| These households are the first part of the sample. To include also households not listed |  |
| in the telephone book the interviewer takes each "telephone household" as starting |  |
| point to visit the fifth household after the start household (according to a specified rule |  |
| for random route). The households found with that method are the second part of the |  |
| sample. |  |
| Stage 3: Within all households the target persons are drawn with the Next-Birthday- |  |
| Method. |  |$\quad$| The allocation and selection of clusters was done at ZUMA. |
| :--- |
| Remark |


| Design effects | $\mathrm{DEFF}_{\mathrm{c}}=1+(7-1) * 0.02 \approx 1.1$ <br> $\mathrm{DEFF}_{\mathrm{p} 1}=1.25, \mathrm{DEFF}_{\mathrm{p} 2}=1.1$ <br> $\mathrm{DEFF}=1.25 * 1.1 * 1.1 \approx 1.5$ |
| :--- | :--- |
| Remark | $\mathrm{DEFF}_{\mathrm{p} 1}$ is due to unequal selection probabilities within households. Note, that $\mathrm{DEFF}_{\mathrm{p} 1}$ <br> is computed on the basis of data for the whole population (not $15+$ ). $\mathrm{DEFF}_{\mathrm{p} 2}$ is due to <br> unequal selection probabilities of listed and unlisted households. A DEFF <br> p for unequal <br> selection probabilities of households in general is also included. |


| Target response rate | A response rate of $65 \%$ seems to be reasonable. |
| :--- | :--- |
| Problems | The target response rate of $70 \%$ will (probably) not be achieved. |


| Sample size | A net sample of size $n_{\text {net }}=n_{\text {eff }} *$ DEFF $=1,500 * 1.5=2,250$ interviews will be <br> conducted. With $10 \%$ of ineligibles and a response rate of $65 \%$ the gross sample size <br> must be: $2,250 /(0.90 * 0.65)=3,847$. For each cluster 12 individuals have to be <br> drawn. With $10 \%$ ineligibles and a response rate of $65 \%$ that results in an average of 7 <br> interviews per cluster. |
| :--- | :--- |

During the interviews it has to be asked whether a household is listed in the telephone book or not. These data have to be used for weighting adjustments.

| Country: | Belgium |
| :--- | :--- |
| NC: | Geert Loosveldt |
| Survey Institute: | Institute for Social and Political Opinion Research |
| Expert: | Sabine Häder |


| Target Population, <br> Population coverage | Persons 15 years or older who are resident within private households in Belgium, <br> regardless of nationality and citizenship, language or legal status. Homeless and <br> institutional populations are excluded. |
| :--- | :--- |


| Sampling frame | Frame of individuals: <br> In Belgium there is a National Register. The coverage and the updating of the register <br> can be considered as excellent. |
| :--- | :--- |


| Sampling design | Stratified two stage probability sampling <br> Regional stratification: 10 provinces and Brussels <br> Stage 1: The primary sampling units (PSUs) are clusters located in municipalities. The <br> number of clusters for each province is proportional to the size of the population in |
| :--- | :--- |
| each province. For that a list of municipalities with a population distribution ( +15 |  |
| years) for each province is used. The number of clusters in a municipality is |  |
| proportional to the size of its population. |  |
| Stage 2: In each of the 324 clusters in 202 municipalities 10 individuals are selected for <br> the gross sample. The clusters within the municipalities are 'virtual', i.e. they are not <br> regionally defined. That means for each cluster 10 individuals within a municipality are <br> selected from the register by simple random sampling. |  |


| Design effects | The intra class correlation coefficient proposed in the specifications for participating <br> countries $(\rho=0.02)$ is used. <br> $\mathrm{DEFF}_{\mathrm{c}}=1+(6.2-1) * 0.02=1.10$ <br> $\mathrm{DEFF}_{\mathrm{p}}=1$ <br> $\mathrm{DEFF}=1.10$ |
| :--- | :--- |
| Remark | The mean cluster size is computed as number of respondents divided by number of <br> communities, i.e. $2,000 / 202=9.9 \approx 10$. The clusters within the communities are not <br> taken into account, because they are only virtual, that means they do not constitute a <br> residential area. |


| Target response rate | A response rate of $65 \%$ seems to be reasonable. |
| :--- | :--- |
| Problems | The target response rate of $70 \%$ will (probably) not be achieved. |


| Sample size | A net sample of 2,000 interviews will be conducted. With a response rate of $65 \%$ and <br> $5 \%$ of ineligibles the gross sample size must be: <br> $2,000 /(0.65 * 0.95) \approx 3,239 . ~ T h e ~ e f f e c t i v e ~ s a m p l e ~ s i z e ~ i s ~$ |
| :--- | :--- |
| $2,000 / 1.10=1,818$. |  |

Country:
NC:
Survey Institute:
Expert:

## Czech Republic

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STEM
Susan Purdon / Peter Lynn

| Target Population, <br> Population coverage | All persons aged 15 and over living in households in Czech Republic |
| :--- | :--- |


| Sampling frame | The "SIPO" database of households. This is compiled by merging utility lists of <br> households that subscribe to electricity, gas, radio, television or telephone. It is <br> estimated that $98 \%$ of households are on SIPO. The database is managed by a private <br> company, VAKUS, and updated monthly. There is no official population register or <br> register of addresses available for academic purposes in Czech Republic. |
| :--- | :--- |


| Sampling design | Stratified two stage probability sampling <br> Stratum 1: 100 large towns/cities, accounting for 52\% of population according to 2001 <br> Census. Divided into regional sub-strata. In each sub-stratum, households are ordered <br> by postal code and a systematic random sample selected using a fixed interval. For <br> eych household selected in this way, the 2 |
| :--- | :--- |
| lid $4^{\text {th }}, 6^{\text {th }}$ and $8^{\text {th }}$ following households on the |  |
| list are also included, thus making a cluster (PSU) of 5 selected households. In total, |  |
| 374 clusters (1,870 households) will be selected in stratum 1. |  |
| Stratum 2: 6,140 localities, 48\% of population. Divided in to 28 sub-strata, defined by |  |
| NUTS2 region and population size of locality. In each sub-stratum, clusters of 5 |  |
| households will be selected systematically, exactly as in stratum 1. 292 clusters (1,460 |  |
| households) selected. |  |
| The sampling interval will be the same in all sub-strata (both main strata), except for <br> some small variation to anticipate likely differences in response rate (e.g. slightly <br> smaller interval in the largest towns). |  |
| At each household, one person aged 15+ will be selected using a Kish grid procedure. |  |


| Design effects | $\mathrm{DEFF}_{\mathrm{P}}=1.16$, based on estimate of household size distribution from Census 2001. <br> $\mathrm{DEFF}_{\mathrm{c}}=1+(3.15-1) * 0.05=1.1$ <br> $\mathrm{DEFF}=1.16 * 1.11=1.29$ |
| :--- | :--- |
| Remarks | No evidence on roh. 0.05 is a deliberately pessimistic guess, reflecting the "dense" <br> clusters. Even with this pessimism, DEFF is still acceptably small. |


| Target response rate | $70 \%$ |
| :--- | :--- |
| Remarks | Believed to be possible, though $65 \%$ is more common for this kind of sample design. |


| Sample size | Gross sample size $\mathrm{n}_{\text {gross }}=3,330$ <br> With $10 \%$ ineligibles, $70 \%$ response: <br> Net sample size $\mathrm{n}_{\text {net }}=2,10$ interviews <br> Effective sample size $\mathrm{n}_{\text {eff }}=1,630$ |
| :--- | :--- |
| Remarks | Effective sample size should be large enough so long as response rate exceeds $65 \%$ |

Special Features of the
design
Disproportionate sampling by strata to reflect anticipated variation in response rates.

| Country: | Denmark |
| :--- | :--- |
| NC: | Torben Fridberg |
| Survey Institute: | SFI Survey |
| Expert: | Peter Lynn |

Target Population, Persons 15 years or older who are resident in Denmark, regardless of nationality and Population coverage citizenship, language or legal status.

| Sampling frame | Danish Central Person Register (CPR). The CPR has approximately 99.9\% coverage of <br> persons resident in Denmark. All persons who expect to stay in Denmark for at least 3 <br> months are included. Homeless persons without an address are excluded. |
| :--- | :--- |
| Remarks | It is not yet clear whether persons in institutions will be excluded - it is possible that <br> prisons will be the only institutions to be excluded. However, this is not important as <br> the sample is large enough that persons in institutions could be excluded at the analysis <br> stage if needed (they will be few in number and identifiable). |


| Sampling design | Simple random sample of persons born before 1-9-1987 |
| :--- | :--- |


| Design effects | No stratification, no clustering, equal probabilities, so $\mathrm{DEFF}=1.00$ |
| :--- | :--- |


| Target response rate | $75 \%$ is a realistic expectation, based upon other household surveys in Denmark. |
| :--- | :--- |


| Sample size | Gross sample size $\mathrm{n}_{\text {gross }}=3,000$ <br> $75 \%$ response rate <br> Net sample size $\mathrm{n}_{\text {net }}=2,250$ |
| :--- | :--- |

Country:
NC :
Survey Institute:
Expert:

Finland
Heikki Ervasti
Statistics Finland
Peter Lynn

Target Population,
Population coverage

All residents aged 15 and over, excluding only those in institutions (defined as prisoners plus those in permanent care in hospitals and similar institutions), commercial sailors, those temporarily away from their usual address (a small group mainly people studying or working abroad) and those with no valid address information on the register. The total number of exclusions from all these categories is estimated at 70,000 (cf. Population of nearly 4 million). Foreign citizens are included if they have residency status.

| Sampling frame | Population register |
| :--- | :--- |


| Sampling design | Single stage equal probability systematic sample (no clustering). Implicit stratification <br> by region, sex and age. |
| :--- | :--- |


| Design effects | DEFF $=1.0$ (no clustering, equal probabilities). |
| :--- | :--- |


| Target response rate | $75 \%$ |
| :--- | :--- |
| Problems | It is thought that this is realistic |


| Sample size | With a gross sample size $n_{\text {gross }}=2,800,1.5 \%$ ineligibles and a response rate of $75 \%$ a <br> net sample size $n_{\text {net }}=2,800 * 0.985 * 0.75>2,050$ can be expected. |
| :--- | :--- |

Country:
NC :
Survey Institute:
Expert:

France
Bruno Cautres, Etienne Schweisguth, Nadine Mandran
ISL
Siegfried Gabler / Sabine Häder

Target Population, $\quad$ Persons 15 years or older who are resident within private households in France, Population coverage regardless of nationality and citizenship, language or legal status.

| Sampling frame | For stratification: INSEE population data (15 years or older) from the last census 1999 |
| :--- | :--- |
| Remarks | Foreigners are excluded in the INSEE population data |


| Sampling design | Stratified three stage probability sampling <br> Stratification: Table with 9 ZEAT areas and 6 agglomeration classes, i.e. 54 cells (9 <br> empty). Each cell contains the corresponding population size. <br> Stage 1: Allocation of 125 Primary Sampling Units to the defined cells using Cox <br> Method of controlled rounding (done by ZUMA). Selection of PSUs (communities) <br> from the cells according to the allocation. <br> Stage 2: Selection of a fixed number of households from the selected PSUs via a <br> random route procedure. Four start addresses per PSU are selected from the telephone <br> book (no interview in these starting points). Thus, five more households have to be <br> found via random route to get 20 households in the PSU. The households are listed in <br> advance. <br> Stage 3: Selection of an individual within a household via Last-Birthday-Method. An <br> average of 12 interviews will be conducted in each PSU. |
| :--- | :--- |


| Design effects | $\mathrm{DEFF}_{\mathrm{c}}=1+(12-1) * 0.02=1.22$ <br> $\mathrm{DEFF}_{\mathrm{p}}=1.3$ <br> $\mathrm{DEFF}=1.58$ |
| :--- | :--- |
| Remark | $\mathrm{DEFF}_{\mathrm{p}}$ is due to unequal selection probabilities within households. Note, that $\mathrm{DEFF}_{\mathrm{p}}$ is <br> computed on the basis of data for the whole population (not $15+$ ). |


| Target response rate | It is extremely difficult to estimate the response rate since a survey like this with strict <br> probability sampling has been conducted in France only very rarely. The response rate <br> is estimated to be about $60 \%$. |
| :--- | :--- |
| Problems | The target response rate of $70 \%$ will probably not be achieved. |


| Sample size | Effective sample size $n_{\text {eff }}=950$ <br> Net sample size $n_{\text {net }}=950 * 1.58 \approx 1,500$. <br> Gross sample size $\mathrm{n}_{\text {gross }}: 1,500 / 0.6=2,500$. <br> This means, in each registry about 20 households have to be drawn for the gross <br> sample and 12 individuals have to be interviewed. |
| :--- | :--- |
| Problems | The effective sample size as well as the net sample size is below the ESS requirements. <br> This is because of funding restrictions. |


| Country: | Germany |
| :--- | :--- |
| NC: | Jan van Deth, University of Mannheim |
|  | Katja Neller, University of Stuttgart |
| Survey Institute: | Infas |
| Expert: | Seppo Laaksonen |
| Reference Survey: | ALLBUS 2002 |

Target Population, Population coverage

All persons aged 15 and over who are resident in Germany; two independent target populations, one for West Germany incl. West Berlin, and the other for East Germany incl. East Berlin.

| Sampling frame | For stratification and selection of communities: <br> Data base provided by the German Statistical Office (2002) and the Federal Statistical <br> Offices. <br> For selection of individuals: <br> The registers of local residents' registration offices will be used for the selection of the <br> individuals. In these offices, all foreigners living in Germany as well as German <br> citizens are listed. The registers are updated continuously. |
| :--- | :--- |


| Sampling design | Stratified two-stage probability sampling separately for East and West Germany <br> Stratification: districts * regional size categories; 1,085 strata in West Germany and <br> 435 strata in East Germany. <br> Stage 1: Selection of 100 communities (clusters) for West Germany, and 50 for East <br> Germany. The communities are selected with probability proportional to the population <br> size of the community (aged 15+). The number of communities selected from each <br> stratum is determined by a controlled rounding procedure. The number of sample <br> points is 108 in the West, and 55 in the East (some larger communities have more than <br> one sample point). <br> Stage 2: In each sample point an equal size of individuals will be selected by a <br> systematic random selection process. This is done by using the local registers of <br> residents' registration offices. |
| :--- | :--- |


| Design effects | DEFF is derived from the first stage of the sampling design, that means because of <br> clustering, and from differing selection probabilities because of oversampling in East <br> Germany. Given a mean cluster size of $\mathrm{b}=20$ and an intraclass correlation coefficient <br> of $\rho=0.02$ we get <br> $\mathrm{DEFF}_{\mathrm{c}}=1+(20-1) * 0.02=1.38$ <br> $\mathrm{DEFF}_{\mathrm{p}} \approx 1.1 ; \mathrm{DEFF}=\mathrm{DEFF}_{\mathrm{c}} * \mathrm{DEFF}_{\mathrm{p}}=1.52$ |
| :--- | :--- |


| Target response rate | $70 \%$ |
| :--- | :--- |
| Problems | A response rate of $70 \%$ seems to be rather unrealistic. In surveys with a comparable <br> content and the same design about $50 \%$ are usual. |


| Sample size | Gross sample size $\mathrm{n}_{\text {gross }}=4,868$, Ineligibles: $10 \%$ <br> Net sample size $\mathrm{n}_{\text {net }}=3,066$ persons $(1,020$ for the East and 2,046 for the West $)$ <br> The target sample size is so big that there will not be any problem to achieve the ESS <br> effective sample size $\mathrm{n}_{\text {eff }}(3,066 / 1.52=2,017>1,500)$. |
| :--- | :--- |


| Special Features of the <br> design | -Generally, the advantage of sampling from address registers is that at least <br> minimal information is available on non-response cases (such as age, sex). For <br> these variables the extent of sample bias can be estimated by matching participants <br> and non-participants. Infas also offers selectivity analyses using information on <br> individuals. If required, Infas can calculate redressment/adjustment weights, <br> otherwise self-weighting is used. <br> Oversampling of the East German population |
| :--- | :--- |

Country:
NC :
Survey Institute:
Expert:

## Greece

Yannis Voulgaris (y_v@ekke.gr)
EKKE (Greek National Centre for Social Research)
Peter Lynn

Target Population, Population coverage

All persons aged 15 and over living in private households in Greece, excluding the Cyclades islands and the Dodecanese islands, apart from Rhodes. The "homeless" and institutionalised population are excluded.

| Sampling frame | Area-based sampling, using frame of area units constructed from 2001 Greek Census. |
| :--- | :--- |

## Sampling design

Stratified three-stage probability sampling
stages: area units (average 40 households), households, persons.

Stage 1: Area units (PSUs) are sorted into 101 strata. Greater Athens is divided into 31 geographical strata, Greater Salonica into 9, and the rest of Greece into 61 strata, defined by degree of urbanisation (up to 8 categories) and region ( 10 regions). Sample size is allocated to strata in proportion to the (Census) number of households. The sample size is then divided into PSUs, based on 6,7 , or 8 sample households per PSU (fixed within strata). Within each stratum, PSUs are selected PPS. Total number of sampled PSUs is 438.

Stage 2: Within each sampled area unit, interviewers will make a complete listing of all resident households (dwellings/ doors). For all towns and cities ( $83 \%$ of PSUs), the interviewer will be given a Census map clearly showing the area unit; for rural areas field supervisors will create a rough map and description of the boundaries. The completed listing will be passed to a field supervisor, who will then apply a random start and interval to select households systematically.

Stage 3: 1 resident ( $15+$ ) selected at random using Kish grid.

| Design effects | $\mathrm{DEFF}_{\mathrm{p}}=1.18$ (guess); $\mathrm{DEFF}_{\text {str }}=0.99$ (guess); <br> $\mathrm{DEFF}_{\mathrm{c}}=1.15$, based on $\rho=0.04, \mathrm{~b}=4.8$ <br> $\mathrm{DEFF}=1.18 * 0.99 * 1.15=1.34$ |
| :--- | :--- |
| Problems | None. These are guesses, but sample size assumes DEFF will not exceed 1.4. |


| Target response rate | $70 \%$ |
| :--- | :--- |
| Problems | This is really difficult to predict as there has been no similar survey in Greece |


| Sample size | Gross sample size $\mathrm{n}_{\text {gross }}=3,100$ <br> $68 \%$ response rate <br>  <br>  <br>  <br>  <br> Net sample size $\mathrm{n}_{\text {net }}=2,100$ interviews <br> Effective sample size $\mathrm{n}_{\text {eff }}=1,570$ if DEFF $=1.34$ $\mathbf{l}$ |
| :--- | :--- |

Special Features of the design

The design produces an equal-probability sample of households. The only variation in selection probabilities - and hence weighting - will be due to selection of a random person within households.
Country: Hungary

NC: Peter Robert
Survey Institute:
Tarki
Expert:
Seppo Laaksonen
Reference Survey: ISSP surveys carried out between 1986 and 2002

| Target Population, <br> Population coverage | All persons aged 15 and over who have a status of permanent residents in the territory <br> of Hungary |
| :--- | :--- |


| Sampling frame | For selection of individuals: <br> The logs of Central Register and Election Office are used. The Central Register and <br> Election Office data are fully up-to-date, they can be used as a high quality source for <br> sampling purposes. The database is updated online according to any kind of change in <br> the population records (births, deaths, migration). All residents in Hungary are <br> recorded and followed reliably and the database is accessible for sampling purposes on <br> a legal way. |
| :--- | :--- |
| Remarks | Some groups of people are hardly accessible with the help of the residents' database, <br> such as homeless people, people living temporarily abroad, people with uncertain <br> residential background |


| Sampling design | Stratified two-stage probability sample: <br> Stage 1: Settlements are categorized by size, geographical area and administrative <br> status in different groups. The selection of 143 settlements (PSUs) is done Proportional <br> to population size. <br> Stage 2: For each settlement a number of individuals is specified taking into account <br> the population size relative to the whole target population. The target sample units <br> (individuals meeting the age criteria, i.e. being older than 15 years) are selected <br> randomly from the pre-selected register database. The pre-selection of the register <br> database follows the specified age criteria. According to the above described sampling <br> procedure it is assured that each person being resident in Hungary and meeting the <br> specified age criteria has equal likelihood to be chosen. |
| :--- | :--- |


| Design effects | $\mathrm{DEFF}_{\mathrm{p}}=1$ <br> $\mathrm{DEFF}_{\mathrm{c}}=1+(10.5-1) * 0.02=1.19$ <br> $\mathrm{DEFF}=1.19$ |
| :--- | :--- |


| Target response rate | $65 \%$ |
| :--- | :--- |
| Remarks | In earlier similar surveys the response rate has been between 55-65\%. |


| Sample size | $n_{\text {gross }}=2,450$, ineligibles $6.5 \%(\mathrm{n}=160)$, response rate $65 \%$ <br> $\mathrm{n}_{\text {net }}=1,500 \mathrm{n}_{\text {eff }}=1,500 / 1.19=1260$ |
| :--- | :--- |
| Remarks | Both the net sample size and the effective sample size are below the ESS requirements. |

Special Features of the design

There is no special oversampling or weighting technique used during the sampling procedure, however, after the data processing, there is a weight calculated for the whole sample based on the four-dimensional distribution of the demographic variables: gender, age, school, type of settlement. This weight provides an adjustment to the distribution of the target population.

Country:
NC:
Survey Institute:
Expert:
Reference study:

## Ireland

Richard Sinnott
Economic and Social Research Institute (ESRI)
Sabine Häder
Irish Social and Political Attitudes Survey 2002 and others

| Target Population, <br> Population coverage | Persons aged 15 years and over who are resident in private households in the Republic <br> of Ireland |
| :--- | :--- |
| Problems | 15 year olds will only be interviewed with parental consent |


| Sampling frame | For stratification and selection of addresses: <br> Computer-based National Electoral Register. This register is updated annually. The <br> version of $2002 / 03$ will be used. |
| :--- | :--- |


| Sampling design | Stratified three-stage probability sampling <br> PSUs are aggregates of District Electoral Divisions (DEDs). There is a total of 3,440 <br> DEDs in Ireland. <br> PSUs are ordered geographically (north-south, east-west). This makes it possible to <br> take a systematic sample giving an implicit stratification. |
| :--- | :--- |

Stage 1: Selection of 220 PSUs proportional to size of population.
Stage 2: Systematic sample of 19 to 20 addresses within each PSU. Addresses are sorted by geographical propinquity.

Stage 3: Random selection of one individual within each household (Next-BirthdayMethod)

- Given the nature of the sampling frame, which is the only one available in Ireland, one is constrained to use the population aged 18 and over in selecting the PSUs. The sample from the electoral register is a sample of addresses. In effect this is very close to a sample of households in Ireland as there is a very low level of multi-household occupancy at addresses in Ireland. If there is more than one household behind an address the household to be interviewed is selected via Kish grid (given the nature of the electoral register the chances of more than one household at any address are very small indeed).

| Design effects | $\mathrm{DEFF}_{\mathrm{c}}=1+(11-1) * 0.02=1.2$ <br> $\mathrm{DEFF}_{\mathrm{p}}=1.33$ <br> $\mathrm{DEFF}=1.6$ |
| :--- | :--- |
| Remark | Differential household sizes between PSUs are already included in the calculation of <br> the design effect DEFF $_{\mathrm{p}}$. |


| Target response rate | About $63 \%$ |
| :--- | :--- |
| Problems | The target response rate (70\%) will not been reached. |


| Sample size | A net sample of size $n_{\text {net }}=n_{\text {eff }} *$ DEFF $=1,500 * 1.6=2,400$ interviews will be <br> conducted. With $10 \%$ of ineligibles and a response rate of about $63 \%$ the gross sample <br> size must be $n_{\text {gross }}=2,400 /(0.63 * 0.9)=4,233$. For each cluster 19 to 20 addresses <br> have to be drawn. That results in an average of 11 interviews per cluster. |
| :--- | :--- |

Special Features of the design

Given the sample design it is necessary to re-weight the data according to household size due to unequal selection probabilities. This is implemented as a standard procedure in our samples and is discussed in the written documentation previously provided by the ESRI.

Country:
NC :
Survey Institute:
Expert:

## Israel

Noah Lewin-Epstein (noah1@post.tau.ac.il)
Institute for Social Research/Cohen Institute
Siegfried Gabler

| Target Population, <br> Population coverage | The survey will be representative of all persons aged 15 and over (no upper age limit) <br> resident within private households in Israel, regardless of their nationality, citizenship, <br> language or legal status. <br> A) The target population includes the Jewish population residing in West-Bank and <br> the Gaza-Strip (about 200,000). |
| :--- | :--- |
| Remarks | B)The target population does not include the Palestinian residents of (East) <br> Jerusalem (about 200,000). <br> The size of the target population is 3,835,994. |
| Concerning A) <br> This is a deviation from the ESS-specifications but „these individuals are very much a <br> part of the fabric of Israeli society". <br> Concerning B) <br> "This is mostly for practical and technical reasons. These residents refuse to cooperate <br> with Israeli official or semi-official activities. The Israeli Central Bureau of Statistics <br> (CBS) has noted the extreme difficulties in collecting data in these areas and has <br> essentially given up on it. Moreover, in the current situation interviewers would be at <br> risk walking in these neighbourhoods and searching for addresses." |  |


| Sampling frame | For selection of households: <br> The sampling frame is the telephone directory (containing 95\% of all households; <br> matching the list of addresses is possible) |
| :--- | :--- |
| Problems | Population of <br> $-\quad$ nomads (Beduin), not included in the sampling frame <br> $-\quad$ Arabs in small rural communities (1/3 do not have a phone) |


| Sampling design | Stratified three stage probability sampling: <br> Israel is stratified into 11 strata consisting of 2,356 statistical areas (PSUs). <br> Stage 1: Statistical areas <br> 200 PSUs are selected proportional to sizes (all persons aged 15 and over within a <br> statistical area) <br> Stage 2: Housing units <br> Fixed number of 18 housing units inside a statistical area is selected by simple random <br> sampling. If a housing unit contains more than one household each of them belongs <br> with equal probability to the sample. <br> Stage 3: Individuals <br> A person aged 15 and over in each of the selected households is selected by Last <br> Birthday Method |
| :--- | :--- |


| Design effects | The intra-cluster correlation is assumed to be $\rho=0.02$ <br> $\mathrm{DEFF}_{\mathrm{c}}=1+(12-1) * 0.02=1.22$ <br> $\mathrm{DEFF}_{\mathrm{p} 1}=1.29$ <br> $\mathrm{DEFF}_{\mathrm{p} 2}=1.01($ Oversampling $) \mathrm{DEFF}=\mathrm{DEFF}_{\mathrm{c}} * \mathrm{DEFF}_{\mathrm{p} 1} * \mathrm{DEFF}_{\mathrm{p} 2}=1.59$ <br> RemarkThe sample size in the first stratum is 0.84 of the whole sample size. |
| :--- | :--- |


| Target response rate | $70 \%$ |
| :--- | :--- |


| Sample size | Gross sample size $\mathrm{n}_{\text {gross }}=200 * 18=3,600$ |
| :--- | :--- |
|  | Net sample size $\mathrm{n}_{\text {net }}=($ Gross sample size $-\mathrm{p} \%$ ineligibles $) *$ Target response rate $=$ |
| $3,600 * 0.95 * 0.7=2,394$ |  |
|  | Effective sample size $\mathrm{n}_{\text {eff }}=$ Net sample size $/$ DEFF |
|  | $=3,600 * 0.95 * 0.7 / 1.59=1,506$ |


| Special Features of the <br> design | Oversampling (about 1.4) of Arab-residents |
| :--- | :--- |
| Problems | Oversampling influences the design effect |

Country: Italy

NC: Antonio Schizzerotto, University of Milano Biccocca

Survey Institute:
Expert:

TNS Abacus
Peter Lynn

| Target Population, | All persons aged 15 and over, currently resident at a private address in Italy, regardless |
| :--- | :--- | Population coverage of nationality or legal status.


| Sampling frame | For selection of addresses: <br> Electoral register |
| :--- | :--- |
| Problems | Persons living at an address where there are no registered electors are excluded. The <br> number of such persons is not known, but is believed to be very small. |

## Sampling design

Stratified four stage probability sampling.
Stage 1: Municipalities are divided into 4 strata. Stratum A contains all municipalities with a population of $100,000+$ residents. The other municipalities are divided into 3 regional strata (north-west; north-east and centre; south and islands). In each stratum, the number of selections is proportional to the total population of the strata. Muncipalities are selected with replacement with probability proportional to population size. The number of selections will be 125 , though the number of distinct selected municipalities will be smaller, as some will be selected more than once (the largest municipality, Roma, would be expected to be selected 6 times, for example, and the second-largest, Milano, 3 times).

The number of addresses to select from each municipality, $n_{i}$, is 24 multiplied by the number of times the municipality was selected.

Stage 2: $n_{i} / 6$ electoral precincts are selected (electoral precincts contain an average of 600 electors.) For logistical reasons, it is not possible to select precincts PPS so they will be selected with equal probabilities. However, there is very little variance in size of precincts, and the size will in any case be captured for possible use in weighting if needed.

Stage 3: 6 addresses are selected from each electoral precinct with probability proportional to the number of electors at the address. This is achieved by selecting a simple random sample of electors.

Stage 4: The interviewer will list all persons aged $15+$ currently resident at the address and will randomly select one for interview using a Kish-grid method.

The first stage will be carried out by the National Co-ordinator; the second, third and fourth stages will be carried out by the fieldwork contractor.

| Design effects | There is some small variation in selection probabilities to be expected, due to dif- <br> ferences between number of electors at an address and number of residents aged 15+: $^{D_{\mathrm{p}}=1.01}$ <br> $\mathrm{DEFF}_{\mathrm{c}}=1+(4-1) * 0.03=1.09$ <br> $\mathrm{DEFF}^{\prime}=1.10$ |
| :--- | :--- |


| Target response rate | $70 \%$ |
| :--- | :--- |
| Problems | Realistically, high-60s are possible. Sample size calculation is based on assumption of <br> $68 \%$ response. |


| Sample size | $\mathrm{n}_{\text {gross }}=3,000 ;$ anticipated $98 \%$ eligible, $68 \%$ response, <br> $\mathrm{n}_{\text {net }}=2,000$ <br> $\mathrm{n}_{\text {eff }}=1,820$ |
| :--- | :--- |


| Special Features of the <br> design | Weighting by $N_{j} / E_{j}$, where $N_{j}$ is number of current residents aged $15+$ at address $j$ and <br> $E_{j}$ is number of registered electors at address $j$. Note that both of these data will be <br> collected by interviewers, as it is not possible to collect $E_{j}$ at the time of sampling from <br> the electoral registers. |
| :--- | :--- |

Country:
NC :
Survey Institute
Expert:

## Luxembourg

$$
\begin{aligned}
& \text { Uwe Warner, CEPS (uwe.warner@ceps.lu) } \\
& \text { CEPS } \\
& \text { Peter Lynn }
\end{aligned}
$$

| $\begin{array}{l}\text { Target Population, } \\ \text { Population coverage }\end{array}$ | All persons aged 15 and over living in private households in Luxembourg. |
| :--- | :--- |


| Sampling frame | Social security register (Inspection générale de la sécurite sociale (IGSS), as at <br> December 2001). This is a list of "tax units" (one or more persons who are treated <br> collectively for tax purposes). Approximately 91\% of resident adults are on the IGSS, <br> the exceptions being EU civil servants and employees of international organizations <br> and foreign banks. |
| :--- | :--- |
| Problems | Persons living in households where no household member is on the IGSS have no <br> chance of selection. This under-coverage of the target population, will consist mostly <br> of the categories mentioned above. It will also include some persons who were in the <br> same household as a tax unit head in December 2001, but not at the time of ESS <br> fieldwork. This could include some young persons who have become tax registered for <br> the first time in this period and some members of couples who have "split" during this <br> period. |


| Sampling design | Stage 1: The IGSS is sorted into 41 strata, defined by labour market status (14 <br> categories) x size of household (3 categories) (there is one empty stratum). An <br> independent random sample of tax units is selected from each stratum, using variable <br> sampling fractions in order to achieve a minimum selected sample of 60 units per <br> stratum (with one exception, a particularly small stratum where only 11 units are to be <br> selected). The largest sampling fraction is 20 times the smallest one, and the DEFF due <br> to variable sampling fractions at this stage is estimated to be 1.33. In total, 5,033 tax <br> units will be selected. The sample will not be clustered. <br> Stage 2: The household to which each of the 5,033 tax unit heads belongs will be <br> identified in the field and a random selection of one person aged 15 or over in the <br> household made using the last birthday method. |
| :--- | :--- |
| Remark | The selection probability for each person in household $i$ in stratum $j$ is proportional to <br> $F_{j} R_{i j} / N_{i j}$, where $R_{i j}$ is the number of tax reference persons in the household, $N_{i j}$ is the <br> total number of eligible persons in the household, and $F_{j}$ is the sampling fraction of tax <br> units in stratum $j$. Both $R_{i j}$ and $N_{i j}$ will be collected in the ESS interview and included <br> on the data set, to permit calculation of design weights. |


| Design effects | $\mathrm{DEFF}_{\mathrm{P}}=1.40$ (guess); <br> DEFF <br> $\mathrm{DEFF}=1.00$ (no clustering); <br> Remark |
| :--- | :--- |
| DEFF is difficult to predict, as the distribution of $R_{i j} / N_{i j}$ is unknown. If $R_{i j} / N_{i j}=1.0$ <br> in all cases, then DEFF can be expected to be 1.33. Extra variation in $R_{i j} / N_{i j}$ is likely <br> to increase DEFF above this figure. DEFF could easily be as high as 1.60 if most <br> households have $R_{i j}=1$. |  |


| Target response rate | $30 \%$ |
| :--- | :--- |
| Problems | The target response rate will not be reached. |


| Sample size | Gross sample size $\mathrm{n}_{\text {gross }}=5,033 ; 30 \%$ response: <br> Net sample size $\mathrm{n}_{\text {net }}=1,510$ interviews <br> Effective sample size $\mathrm{n}_{\text {eff }}=1,080$ if DEFF $=1.40$ |
| :--- | :--- |
|  |  |

[^3]
## Country: The Netherlands

NC: Peer Scheepers
Survey Institute:
Expert:

GfK
Siegfried Gabler

Target Population, Population coverage

The survey will be representative of all persons aged 15 and over (no upper age limit) resident within private households in the Netherlands, regardless of their nationality, citizenship, language or legal status. Intra-murals (1.3\%) and the sailing and trucking persons are excluded from the gross sample.

| Sampling frame | For selection of addresses: <br> Postal delivery points from 'PTT-afgiftenpuntenbestand' |
| :--- | :--- |


| Sampling design | Two stage probability sampling: (no stratification) |
| :--- | :--- |
| Stage 1: Postal delivery points (excluding P.O. boxes and business addresses) selected |  |
| with equal probabilities |  |
| Stage 2: Person within a household are selected with Next-Birthday- Method |  |
| lf more than one household belongs to a postal delivery point up to 5 households are |  |
| added to the gross sample and one person within a household is selected (Next- |  |
| Birthday-Method). |  |


| Design effects | The sampling design does not contain any clustering. The design effect is only due to <br> differing selection probabilities. <br> $\mathrm{DEFF}_{\mathrm{p}}=1.19$ |
| :--- | :--- |


| Target response rate | $70 \%$ |
| :--- | :--- |


| Sample size | Gross sample size $n_{\text {gross }}=3,565$ persons |
| :--- | :--- |
|  | Net sample size $n_{\text {net }}=($ Gross sample size $-5 \%$ ineligibles $) *$ Target response rate $=$ |
| 2,371 |  |
|  | Effective sample size $n_{\text {eff }}=n_{\text {net }} /$ DEFF $_{\mathrm{P}} \approx 2,000$ |


| Country: | Norway |
| :--- | :--- |
| NC: | Kristen Ringdal (kristen.Ringdal@svt.ntnu.no) |
| Survey Institute: | Statistics Norway (Oyven Kleven = kle@ssb.no) |
| Expert: | Seppo Laaksonen |
| Reference Survey: | The Norwegian Electoral Survey |


| Target Population, <br> Population coverage | All persons aged 15 and over living in Norway. |
| :--- | :--- |


| Sampling frame | For selection of individuals: <br> The BEBAS Population Register which is a working copy of the National Population <br> Register from 1 <br> st January 1999. This will be updated monthly. For the ESS, the latest <br> update will be available from July 2002. The frame covers Norwegian citizens who are <br> not registered living in another country and non-Norwegian citizens who are registered <br> living in Norway excluding students. |
| :--- | :--- |


| Sampling design | Stratified two-stage probability sampling <br> Stage 1: The country is divided into a set of PSUs. The starting point are the 435 <br> municipalities in 19 counties. Some municipalities are collapsed and thus 363 PSUs are <br> obtained. These are divided into 109 strata. All PSUs with more than 30,000 <br> inhabitants, and some with a population number between 25,000 and 30,000, constitute <br> separate strata. For the remaining PSUs as homogeneous strata as possible are formed. <br> As stratification variables are used: industrial structure, number of inhabitants, <br> centrality, communication structures, commuting patterns, trade areas and local media <br> coverage. The stratification is done in such a way that no PSU has less than 7\% of the <br> total population in its stratum, and separately for each county. PSUs that are separate <br> strata (thus, larger municipalities) are drawn with certainty (type I). The remaining are <br> drawn with a probability proportional to the number of inhabitants in the sample area <br> (type II). Thus, one PSU (sample area) from each stratum is selected, together 109 <br> PSUs. <br> Stage 2: A sample of survey units = persons is selected from the 109 strata using |
| :--- | :--- |
| systematic random sampling. The sampling fraction at the second stage is proportional |  |
| to the inverse selection probability at the first stage. For type I, 12 units are selected |  |
| and for type II, 8 units, respectively. It is thus aimed at giving self-weighting when |  |
| both weights are taken into consideration. |  |


| Design effects | Total design effect $\mathrm{DEFF}=1.5$ based on the experience of the previous surveys. For <br> $\rho=0.02, \mathrm{DEFF}_{\mathrm{c}} \approx 1.18, \mathrm{DEFF}_{\mathrm{p}} \approx 1.27$ |
| :--- | :--- |


| Target response rate | $70 \%$ |
| :--- | :--- |


| Sample size | $n_{\text {gross }}=3,215$, for about $10-20$ ineligibles and a response rate of $70 \% n_{\text {net }}=2,250$ <br> $n_{\text {eff }}=n_{\text {net }} / D E F F=1,500$ |
| :--- | :--- |

Country:
NC:
Survey Institute:

Expert:
Reference Survey: Poles-2000, Eurobarometer 38.1

| Target Population, <br> Population coverage | All persons aged 15 and over who have a status of permanent residents in the territory <br> of Poland. |
| :--- | :--- |


| Sampling frame | For selection of individuals: <br> The logs of the Ministry of Internal Affairs and Administration (PESEL) are used. The <br> PESEL data are a high quality source for sampling purposes. It is updated online for <br> births, deaths, and any change of permanent living address. <br> Some categories of people are not covered by PESEL data: institutional population, <br> foreigners working at black market, homeless, people temporary not available. |
| :--- | :--- |


| Sampling design | Stratified one / two - stage probability sampling <br> The country will be divided into two exclusive and exhaustive parts. <br> Stage 1: The first part covers the population living in towns of 100,000 inhabitants or <br> more. This part of a sample will be selected as a simple random sample. <br> The number of this kind of towns is 42, and their target population about $31 \%$ of the <br> whole population. <br> Stage 2: The second part corresponds to the rest of the population - people living in <br> towns of 99,999 inhabitants or less and people living in rural areas. This part of the <br> sample will be stratified and clustered (158 clusters). <br> The sampling of the second part is based on a two-stage design: <br> The sampling frame is first stratified by geography (8 regions) and urbanicity (4 <br> categories), providing 32 strata. For each stratum, the required number of PSUs is <br> established, taking into account a constant cluster size. Then the PSUs are selected with <br> probability proportional to size. The definition of a PSU is different for urban vs. rural <br> areas. For urban areas, a PSU is equivalent to a town, whereas for rural areas, it is <br> equivalent to a village. <br> In the second stage, a cluster of respondents is selected in each PSU. The procedure <br> bases on generating random numbers and checking, if a pointed person meets the <br> criterion of a date of birth. It is assumed that the strategy is equivalent to simple <br> random sampling. A cluster size of 12 individuals seems to be optimal. |
| :--- | :--- |


| Design effects | There is no clustering design effect in the first part of the sample. <br> In the second part, the clustering design effect is anticipated to be $D E F F_{c}=1.22$. Thus, <br> the weighted average for the total sample is $D E F F_{c}=1.14 . T_{c} D E F F_{p}=1.02$ results <br> from unequal probabilities of selection in different urbanicity categories. <br> The average total DEFF $=1.16$. |
| :--- | :--- |


| Target response rate | $70.7 \%$ on average but lower in big cities and higher in rural areas. |
| :--- | :--- |


| Sample size | Gross sample size $\mathrm{n}_{\text {gross }}=2,978$ <br> Overcoverage and ineligible cases $=151$ <br> Net sample size $\mathrm{n}_{\text {net }}=2,000$ persons <br> Effective sample size $\mathrm{n}_{\text {eff }}=1,718$ |
| :--- | :--- |

Special Features of th
The differences in anticipated response rates by region (e.g. in Warszawa just above design $50 \%$ and in some rural areas higher than $80 \%$ ) are taken into account.

## Country: Portugal <br> NC.

Survey Institute:
Expert:
Jorge Vala, University of Lisbon (jmvs@ics.ul.pt)
Euroteste (euroteste@ip.pt)
Reference Survey:
Seppo Laaksonen
European Values Study 1999

Target Population,
Population coverage
All persons aged 15 and over who are resident in the mainland of Portugal (this excludes some islands) and who are not permanently living in institutions or in similar environments. However, there are some population groups which cannot be reached and factually will be excluded from the target population (altogether the proportion of these groups - people living in localities with less than 10 dwellings - is less $3 \%$ ):
people who cannot be interviewed due to language (foreign-based) and ability (mainly elderly people) problems,
some small rural villages which are to be too expensive to interview,
homeless people,
people who are residents of Portugal but working most of time outside the country, retired and other non-employed foreigners mainly in Algarve who's status is not clear and who are living in special residences.

## Sampling frame

## For stratification:

The frame covers all regions of the mainland of Portugal, except some rural villages. The information for the frame is available from the population census of 2001 for all persons 15 years and older.

| Sampling design | Stratified four-stage probability sampling <br> Stratification: <br> $-\quad 5$ regions by gender, <br> $-\quad 3$ age groups by gender, <br> $-\quad 5$ regions $* 5$ municipality (habitat) size classes <br> Altogether 22 strata because the biggest size class is missing in three regions. Within <br> each stratum, the number of the so-called localities (these are administrative areas with <br> varying sizes; e.g. the major part of Lisboa is one locality) is known. For the ESS 100 <br> localities are selected. This number is allocated to each stratum proportionally to size. <br> Respectively, the number of the anticipated respondents for each stratum is determined <br> based on proportional allocation. Moreover, some over-sampling for such strata where <br> the anticipated response rate is lower than the average is required. <br> Stage 1: Selection of 100 localities |
| :--- | :--- |
| Stage 2: Selection of one or more PSUs from each locality proportionally to size (all <br> together 150 PSU's) |  |
| Stage 3: In each PSU street study (random-route technique) for selection of <br> households. The starting point is defined using two different ways: For the localities <br> where a good map is available (this is a typical case), the select the co-ordinates for the <br> starting point with a random method, and for the others localities, they use a random <br> selection with the help of the telephone book. <br> The proportion will be 50\% for each method. In the small localities is more difficult to <br> obtain a good map. <br> Stage 4: Selection of the individuals: Last birthday Method. |  |


| Design effects | ${\text { Average cluster size }=7.5, \text { hence } \mathrm{DEFF}_{\mathrm{c}}=1+(7.5-1) * 0.02=1.13}^{\mathrm{DEFF}}$ <br> $\mathrm{DEFF}=1.1$ |
| :--- | :--- |
| Problems | $\mathrm{DEFF}_{\mathrm{p}}$ seems to be to low |


| Target response rate | $75 \%$ |
| :--- | :--- |
| Problems | The estimation is very optimistic. |


| Sample size | $\mathrm{n}_{\text {gross }}=2,260$ <br> $\mathrm{n}_{\text {net }}=1,695$ <br> $\mathrm{n}_{\text {eff }}=1,367$ |
| :--- | :--- |
| Problems | Does not fulfil the two requirements of the ESS, that is, the effective sample size is <br> below 1,500, and the anticipated net sample below 2,000, respectively. |


| Special Features of the <br> design | Small regional over-sampling due to anticipated response rates |
| :--- | :--- |


| Country: | Slovenia |
| :--- | :--- |
| NC: | Brina Malnar (brina.malnar@Uni-Lj.si) |
| Survey Institute: | Public Opinion and Mass Communication Research Center (CJMMK) |
| Expert: | Siegfried Gabler |

Target Population,
Population coverag Population coverage

The survey will be representative of all persons aged 15 and over (no upper age limit) resident within private households in Slovenia, regardless of their nationality, citizenship, language or legal status. The size of Slovenian population is sligthly below 2 million ( $1,965,000$ ).

| Sampling frame | For selection of individuals: <br> The sampling frame is the Central register of population (CRP) and includes all <br> residents with permanent address, citizens and non-citizens (in principle at least 99\% of <br> the population). Institutionalized persons (army, prisons) are included with their <br> permanent addresses, but are unlikely to be reached by interviewers. The Central <br> register represents a rather 'ideal' sampling frame for survey research. It is regularly <br> updated. A maximum of 10\% ineligible cases is expected, which is a conservative <br> estimate, based on previous surveys. |
| :--- | :--- |


| Sampling design | Stratified two stage probability sampling: <br> Slovenia can be divided into Clusters of Enumeration Areas(CEA). The total number <br> of CEA is about 9,000. CEA are first stratified according to 12 regions $* 6$ types of <br> settlements. <br> Stage 1: Selection of 150 PSUs <br> Selection of fixed numbers of CEA inside strata is made by probability proportional to <br> size of CEA. <br> Stage 2: Selection of 15 SSUs per PSU <br> Fixed number of individuals inside CEA is selected by simple random sampling. |
| :--- | :--- |


| Design effects | The intra-cluster correlation coefficient is estimated as $\rho=0.05$, which is a <br> conservative estimate, based on a 1997 sample with 140 PSUs. <br> DEFF $=$ DEFF $_{\mathrm{c}}=1+(9-1) * 0.05=1.4$ |
| :--- | :--- |


\section*{| Target response rate | $66 \%$ |
| :--- | :--- |}

Sample size $\quad$| $\mathrm{n}_{\text {gross }}$ | $=150 * 15=2,250$ |
| ---: | :--- |
| $\mathrm{n}_{\text {net }}$ | $=\left(\mathrm{n}_{\text {gross }}-10 \%\right.$ ineligibles $) *$ Target response rate |
|  | $=(2,250-225) * 0.66=1,336$ |
| $\mathrm{n}_{\text {eff }}$ | $=\mathrm{n}_{\text {net }} / \mathrm{DEFF}=1,336 / 1.4=955$ |

Country: Spain

NC: Mariano Torcal (mariano.torcal@cpis.upf.es)
Survey Institute: Demoscopia (j.barandiaran@demoscopia.com)
Expert:
Reference Survey:
Seppo Laaksonen and Susan Purdon
EPA (Encuesta de Población Activa). It is the Spanish labor sample, conducted by the Public Statistical Office of Spain (INE)

Target Population,
Population coverage
All persons aged 15 and over who have a status of permanent residents and living in private houses in Spain, except Ceuta and Melilla (the exclusion is based on a practical problem, covering $0.3 \%$ of the Spanish population)

| Sampling frame | For stratification and selection of households: <br> There are two official lists of inhabitants: the national Census conducted every ten <br> years by the INE and the municipal roll continuously updated. INE (Public Statistical <br> Office of Spain) has selected a master sample of electoral sections from this frame. <br> This is a large representative sample with full sociodemographic data from Spain. The <br> master sample is composed of nearly 3,500 sections (from a national total of 33,000) <br> and 65,000 households. The master sample is continuously updated. The ESS sample <br> will be drawn from the 2002 master sample. |
| :--- | :--- |


| Sampling design | Stratified two-step/three-stage probability sampling <br> Step 1: Selection of the master sample <br> Step 2: Selection of the ESS sample from the master sample <br> Stratification: Set up of 34 strata resulting from the crossing of 17 regions with two <br> categories of population sizes (areas with a population less than 500,000 inhabitants <br> (= bracket 1) and over 500,000 inhabitants <br> (= bracket 2)). <br> Stage 1: Random selection of electoral sections in each stratum with probability <br> proportional to the number of households in the section. The total number of sections is <br> $346 ; 121$ of them in urban areas. Sections will be selected by the INE from the INE <br> master sample sections frame. <br> Stage 2: Random selection of 10 (bracket 1) or 12 households (bracket 2) per section <br> depending on the bracket. The households will be selected by the INE. <br> Stage 3: Selection of an individual per household using the last birthday method. |
| :--- | :--- |


| Design effects | $\mathrm{DEFF}_{\mathrm{p}}=1.16$ <br> $\mathrm{DEFF}_{\mathrm{c}}=1.20$ <br> $\mathrm{DEFF}=1.39$ |
| :--- | :--- |
| Target response rate | $70 \%$. |
| Remarks | Very optimistic. For the computation of the sample sizes much lower response rates <br> were used. |


| Sample size | ngross <br> Overcoverage and ineligible cases $=$ very low. <br>  <br>  <br>  <br> $n_{\text {net }}=2,076(1,350$ in bracket 1 and 726 in bracket 2) <br> $n_{\text {eff }} \approx 1,500$ <br> It is expected to get 6 interviews per PSU (electoral section). For bracket 1 this means a <br> response rate of $60 \%$, for bracket 2 this means a response rate of $50 \%$. |
| :--- | :--- |
| Remarks | The quality of the frame is very high and overcoverage and undercoverage are <br> expected to be low. |


| Special Features of the <br> design | There is some oversampling for bracket 2, areas with population 500,000+. |
| :--- | :--- |

Country:
NC:
Survey Institute:
Expert:

Sweden
Stefan Svallfors (stefan.svallfors@soc.umu.se)
Statistics Sweden
Susan Purdon/Siegfried Gabler

Target Population, $\quad$ Persons 15 years or older who are resident in Sweden, regardless of nationality and Population coverage citizenship, language or legal status.

| Sampling frame | For selection of individuals: <br> The frame to be used is the register of the population, which includes all individuals <br> living in Sweden. The frame is extremely well suited for this kind of sampling due to <br> the full coverage of the total population. <br> Updating: The register is updated continuously. <br> Individuals that do not reside on the address stated in the register will be traced through <br> the use of complimentary registers. Individuals residing in the country illegally can, <br> naturally, not be selected, but the proportion of such individuals is small in Sweden and <br> is thus unlikely to cause any substantial bias. <br> The sampling frame includes those living in institutions, since there is no way to filter <br> them. |
| :--- | :--- |
| Remark | First contact in Sweden should be best done by telephone for setting up the interview. <br> The register of the population does not include telephone numbers so the selected <br> individuals will be matched for this after the selection process, both by registers and <br> manually in those cases were registers prove to be non sufficient. This procedure will <br> cause a very small number of individuals (about 200) that cannot be traced (due to e.g. <br> non-registered phone numbers, 1.5\%, and so on). An interviewer will be sent to call on <br> the 200 for which no telephone number exists. In addition, "soft" telephone refusals are <br> followed up with a face-to-face contact. |


| Sampling design | Single stage (without clustering) probability sampling <br> Fully random sample of individuals (equal probability selection) born before <br> 1 September 1987 |
| :--- | :--- |


| Design effects | No stratification, no clustering, equal probabilities, so $\operatorname{DEFF}=1.00$ |
| :--- | :--- |


| Target response rate | $75 \%$ (14\% non contact, $11 \%$ refusals) |
| :--- | :--- |


| Sample size | $\mathrm{n}_{\text {gross }}=3,000$ selected persons; <br> $\mathrm{n}_{\text {net }}=(3,000-70) * 75 \%$ response $=2,198$ interviewed <br> $\mathrm{n}_{\text {eff }}=2,198 /$ DEFF $=2,198$ |
| :--- | :--- |
| Remark | The number of ineligibles (died/emigrated/disappeared/living abroad) is about 40, the <br> number of people living in institutions is estimated as about 30. |

[^4]| Country: | Switzerland |
| :--- | :--- |
| NC: | Dominique Joye |
| Survey Institute: | MIS Trend Lausanne |
| Expert: | Sabine Häder |
| Reference Survey: | Eurobarometer 2001 |

Target Population, Population coverage

Persons 15 years or older who are resident within private households in Switzerland, regardless of nationality and citizenship, language or legal status. Homeless and institutional populations are excluded.

| Sampling frame | For stratification: <br> The telephone numbers are coded geographically according to the "Numero postal <br> d'acheminement" (NPA) which defines a general grid over the country. <br> For the selection of households: <br> The telephone register is the only available frame for national social surveys in <br> Switzerland. <br> The Swiss Statistical Office has the possibility to access to the whole telephone register <br> - that means the non-published numbers are included as well - for public interest <br> surveys. Users of mobile phones are also listed in this complete register if they have a <br> long-term contract and no fixed line. <br> The degree of coverage that can be obtained by using this frame is more than 95\% of <br> all households. |
| :--- | :--- |
| Remarks | Households without telephone or with mobile phones without long term contracts are <br> not included in the frame (about $5 \%$ of the households). Holiday houses are rejected. |

## Sampling design

Stratified three stage probability sampling

Regional stratification: six groups of NPAs
Stage 1: Selection of 220 PSUs (NPAs). The sampling is done with probabilities proportional to the number of households listed in the telephone register.

Stage 2: Simple random sampling of about 30 households in the chosen PSUs

Stage 3: Random selection of a respondent in each household according to Kish grid

| Design effects | $\operatorname{DEFF}_{\mathrm{c}}=1+(10-1) * 0.02=1.18$ <br> $\operatorname{DEFF}_{\mathrm{p}}=1.25$ <br> $\mathrm{DEFF}^{2}=1.47$ |
| :--- | :--- |


| Target response rate | A response rate of about $40 \%$ has to be expected in the case of telephone recruitment, <br> even a little bit lower in the case of direct recruitment. |
| :--- | :--- |
| Problems | The target response rate of $70 \%$ will not be achieved. This expectation results from the <br> Eurobarometer 2001 where also only about $40 \%$ could be reached |


| Sample size | Given the target effective sample size $\left(n_{\text {eff }}=1,500\right)$ and the design effect (DEFF $=$ <br> $1.47)$, the net sample will include at least $1,500 * 1.47=2,205$ respondents for <br> Switzerland. <br> The gross sample size will be $\mathrm{n}_{\text {gross }}=6,660$, approximately three times the expected <br> number of interviews. $14 \%$ ineligibles (3\% non contact, $8 \%$ bad number, no household, <br> etc., $3 \%$ still of holidays houses) are expected. <br> In case of telephone recruitment (170 PSUs) a response rate of $40 \%$ is expected, in the <br> case of face to face recruitment (52 PSUs) a response rate of $35 \%$ is realistic. That <br> means $n_{\text {net }}=2,224$. <br> Case telephone recruitment, 170 clusters: <br> $n_{\text {net1 }}=5,100 * 0.86 * 0.4=1,754$ <br> Case FAF recruitment, 52 clusters <br> $n_{\text {net2 }}=1,560 * 0.86 * 0.35=470$ <br> $n_{\text {net }}=1,754+470=2,224$ |
| :--- | :--- |

Special Features of the design

First contact and selection of the target person will be done via face to face in 52 clusters and telephone in 170 clusters with a visit in case of refusal or non-response. The clusters are assigned to the two contact modes randomly, stratified by region. That is a specific method in order to increase the response rate in Switzerland.

| Country: | Turkey |
| :--- | :--- |
| NC: | Yilmaz Esmer |
| Survey Institute: |  |
| Expert: | Sabine Häder / Siegfried Gabler |

Target Population, $\quad$ Persons 15 years or older who are resident within private households in Turkey, Population coverage regardless of nationality and citizenship, language or legal status.

| Sampling frame | To select the provinces and registries: Population data from the last census 2000. <br> To select the households: Voter registries, updated for last election in November 2002. |
| :--- | :--- |
| Problems | The data used for the selection of the provinces and registries are from the whole <br> population, not only from those who are 15 years or older. However, the <br> correlation between both distributions is about 0.99. |
| Foreigners can only be reached if they live in a household where at least one |  |
| Turkish voter is resident. However, the response rate of foreigners living in Turkey |  |
| is predicted to be extremely low anyway (language problems). To verify this, an |  |
| experiment is conducted (see Special Features Box below) |  |
| People living in households without any voter cannot be reached. The number of |  |
| these households is small enough to be neglected. |  |


| Sampling design | Stratified four stage probability sampling |
| :--- | :--- |
| Stage 1: One part of the sample: 7 Self-representing provinces with a population of |  |
| more than 1.7 Mio inhabitants each. |  |
| Second part of the sample: Remaining provinces |  |
| Implicit stratification of these remaining 73 provinces: 9 regions, Human Development |  |
| Index |  |
| Selection of provinces: 33 provinces are selected with systematic proportional to size |  |
| random sampling |  |
| Thus, altogether 40 out of 81 provinces are considered. |  |
| Stage 2: Selection of voter registries within provinces: 200 PSUs (voter registries) are <br> allocated proportional to population size to the selected 40 provinces (done by ZUMA) <br> as follows: 75 registries are selected within the self-representing provinces and 125 <br> registries are selected within the remaining provinces. <br> Voter registries within the provinces are selected with systematic random sampling. |  |


| Sampling design | Stage 3: Selection of households: From each voter registry a list of households is <br> derived. From that, a fixed number of households is selected randomly. <br> Stage 4: Selection of individuals: The person to be interviewed is selected with a <br> "Lottery-method". To do this, the interviewer prepares cards with the names of all <br> persons living in the household. A person belonging to the household draws one card <br> with the name of the person to be interviewed. About 11 individuals have to be <br> interviewed in each registry. |
| :--- | :--- |
| Remark | With the "Lottery-Method" we hope to increase the acceptance of the within- <br> household-selection. |


| Design effects | $\mathrm{DEFF}_{\mathrm{c}}=1+(11-1) * 0.02=1.2$ <br> $\mathrm{DEFF}_{\mathrm{p}}=1.23$ <br> $\mathrm{DEFF}=1.48$ |
| :--- | :--- |
| Remark | $\mathrm{DEFF}_{\mathrm{p}}$ is due to unequal selection probabilities within households. Note, that $\mathrm{DEFF}_{\mathrm{p}}$ is <br> computed on the basis of data for the whole population (not $15+$ ). |


| Target response rate | The response rate is estimated to be about $60 \%$. |
| :--- | :--- |
| Problems | It is extremely difficult to estimate the response rate since a survey like this (i.e. <br> interview length 90 minutes, strict probability sample) has never been conducted in <br> Turkey before. <br> The target response rate of $70 \%$ will probably not be achieved. |


| Sample size | Effective sample $\operatorname{size} \mathrm{n}_{\text {eff }}=1,500$. <br> Net sample size $\mathrm{n}_{\text {net }}=1,500 * 1.48=2,220$. <br> Gross sample size $\mathrm{n}_{\text {gross }}=2,220 / 0.6=3,700$. <br> This means, in each registry about $18-19$ addresses have to be drawn for the gross <br> sample and 11 individuals have to be interviewed. |
| :--- | :--- |

Special Features of the
design
An experiment will be conducted to explore the co-operation rate of foreigners. For that, in 15 additional registries we do not select households but streets. Within these streets the target households will be identified by a random route. This way it is also possible to reach households with only foreigners. The interviews conducted in these experimental registries will not be included in the official ESS data set.
Country: United Kingdom

NC:
Survey Institute:
Expert:
Reference Survey:

Alison Park (a.park@natcen.ac.uk)
National Centre for Social Research/CACI
Siegfried Gabler
British Social Attitudes (BSA): 2000 Survey

| Target Population, | Persons 15 years or older who are resident in the United Kingdom |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | Population coverage (i.e. England, Wales, Scotland and Northern Ireland), regardless of nationality and citizenship, language or legal status.


| Sampling frame | For selection of addresses: <br> There is a postcode address file (PAF) as sampling frame with over 27 million <br> addresses in 1.71 million Postcodes. PAF is the main Address Database for Royal <br> Mail, containing all known UK Postal Addresses and their associated Postcodes and <br> Delivery Point Suffix information. Daily updated by Royal Mail. |
| :--- | :--- |
| Remark | - It has been estimated by the Office of National Statistics that PAF coverage is $96 \%$ <br> of households and $97 \%$ of individuals. This is the best available sample frame in <br> Britain. <br> Some of the delivery point information in a small part of Northern Ireland is <br> excluded. <br> PAF contains also Jersey, Guernsey, and the Isle of Man. |


| Sampling design | For Great Britain (i.e. England, Wales, Scotland): <br> Stratified three stage probability sampling: <br> The (grouped where these have less than 500 delivery points) sectors will be stratified <br> on the basis of 38 region and 1991 Census information. Within each region, sectors are <br> to be listed in increasing order of population density. |
| :--- | :--- |
| Stage 1: 162 postcode sectors (or grouped sectors) will be selected systematically with <br> probability proportional to delivery point count. |  |
| Stage 2: Selection of 24 delivery points |  |
| Stage 3: one person aged 15+ at the selected address. This will be done by random <br> methods as follows: <br> If there are several occupied dwelling units, one dwelling unit is selected using a <br> Kish grid and computer-generated random numbers; <br> $-\quad$Within the (selected) dwelling unit, one person aged 15+ is selected using a Kish <br> grid and computer-generated random numbers. <br> For Northern Ireland: <br> Unclustered sample of 125 addresses drawn from the Valuation and Land Agency's list <br> of domestic properties (not available in Britain). <br> Remark <br> In Scotland, where multiple dwelling units at the address are common, the selection <br> probabilities of addresses are expanded by the 'Multiple Output Indicator' (MOI) <br> which is available on PAF. Thus, multiple-dwelling unit addresses already have higher <br> selection probabilities. To correct for this we calculate the weights as follows: <br> number of dwelling units $\times$ number of adults in selected household <br> MOI <br> Wherever the MOI is correct, the number of dwelling units and MOI cancel each other <br> out, reducing the size of the weights. |  |


| Design effects | DEFFc $=1.26$ (estimated) <br> DEFFp $=1.23$ (after very slight trimming) <br> DEFF $=1.55$ |
| :--- | :--- |

Target response rate $\quad$ Aim $70 \%$ but calculations for the issued sample are based on $65 \%$ response rate.

| Sample size | $\mathrm{n}_{\text {gross }}=4,013$ addresses <br> $\mathrm{n}_{\text {net }}=4,013 * 0.9 * 0.65=2,348$ interviewed <br> $($ England $83 \%$, Wales $5 \%$, Scotland $9 \%$, Northern Ireland $3 \%)$ <br> $\mathrm{n}_{\text {eff }}=2,348 / 1.55=2,348 / 1.55=1,515$ |
| :--- | :--- |
| Remark | $10 \%$ ineligibles <br> The computation of the effective sample size is not quite correct, since there is a <br> boosted sample in Northern Ireland, but which is small compared to the sample in <br> England + Wales + Scotland |

[^5]
## 5. General assessment

Altogether, almost all countries met the sampling requirements for the ESS. Although some countries never had applied a strict random method before we do not have to state any quota elements or any substitution of individuals or households in the plans we signed off. In so far the ESS samples are really a new quality in cross-cultural research. One exception from the ESS requirements is the target response rate of $70 \%$, that was assessed as not realistic in some countries. Another exception affects the target net sample size. In some countries it was not possible to conduct 2,000 interviews due to budget limitations.
We want to give some remarks concerning details of the work of the sampling expert panel:

- The process of co-operation between the National Co-ordinator and the sampling expert usually started with a short description of the planned design given by the NC. As a reminder for the next round here is fixed what should be included:
- the nature of the sampling units at each stage (e.g. addresses, persons);
- description of the frame(s);
- relative selection probabilities (e.g. equal probabilities or disproportionate);
- any stratification to be used, implicit or explicit;
- selected sample size, expected proportion ineligible, expected response rate; etc.
- prediction of design effect
- The target response rate of $70 \%$ was realistic for most countries. Some countries expected to get between $60 \%$ and $70 \%$. In a few countries this was not possible and an even lower response rate had to be expected. On the other hand, some countries hoped to get $75 \%$. In the meantime most countries have finished their fielding period. From that we know that the anticipated response rates were realistic in most countries.
- The minimum net sample size $(\mathrm{n}=2,000)$ could be reached in most countries.
- For the evaluation of the effective sample sizes ( $\mathrm{n}=1500$ ) further analysis are necessary because the $\mathrm{n}_{\text {eff }}$ depends on the assumptions concerning design effects due to clustering and weighting in each country. In almost all countries there was no experience with taking design effects into account when calculating sample sizes.

For two countries we have already estimated the "true" design effects. They are very similar to the predicted response rates (Gabler/Häder/Lynn 2003).
The National Co-ordinators and the Survey Organisations in most countries were not very familiar with the concept of design effects. Therefore, they needed help from the sampling panel in estimating the different components of DEFF and in calculating the sample sizes.

In the next round this problem will be of less relevance because experiences from the first round can be considered.

- In countries with a Design effect $1 \leq$ DEFF $\leq 1.33$ we had problems to convince the National Co-ordinators that a net sample size of 2,000 interviews is required. In a way, the ESS rule that fixes an effective sample size of 1,500 interviews and a net sample size of 2,000 interviews is not very motivating for improving the designs. But it has to be taken into consideration, that a large sample size is a prerequisite for sensibly analysing subgroups of the sample. For the next round an effective sample size of neff $=1,500$ is fixed in the "Specifications for participating countries". However, because of the above mentioned reason a minimum net sample size is not specified.
- In some countries it is usual practice to make first contacts via telephone. In particular, the Northern countries insisted on that procedure because otherwise the response rate would decrease rapidly. Contacting via telephone is of course no problem with an existing frame of individuals and the possibility of detecting the telephone numbers of the selected individuals. Therefore, we changed the rule that first contacts generally had to be attempted in all countries face to face. In Switzerland a special experiment is conducted as part of the ESS to lighten the question whether it is increasing or decreasing the response rate when first contacts are made via telephone.
- We have a large range in the numbers of PSUs. For example, Germany has only 163, Belgium and Austria have 324, Greece has 438 and so on. From the sampling point of view a possibility to decrease the total survey error is to motivate countries with a low number of PSUs to enlarge them.
- In more countries than we had expected we found regularly updated and more or less complete frames of individuals or households that could be used for sampling. The nature of theses frames is of course that they do not contain people who have an illegal status. The only countries where we theoretically could catch them are those few with random route elements in their designs. In the light of this situation the definition of the target population are revised for the next round (rejection of "regardless of ... legal status").
- In a few countries we had to state that the fieldwork had started before we finally signed off the sampling designs. Of course there were only some minor last questions open for clarification but we want to emphasise again that according to the "Specifications for participating countries" the signing off is the precondition for starting the fieldwork.


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Appendix 1: Table of measures of the designs for all countries (Design effects, response rates, No of PSUs, sample sizes)
Appendix 2: Sampling issues in the "Specifications for participating countries"
Appendix 3: Annex 1 of the "Specifications for participating countries" (Rules for estimating design effects)
Appendix 1: Table of measures of the designs for all countries

|  | Frame | Design effect |  |  | $\begin{gathered} \hline \text { Anticipated } \\ \text { Response } \\ \text { rate (\%) } \\ \hline \end{gathered}$ | $\mathbf{n}_{\text {net }}$ | $\mathbf{n g r o s s}$ | $\mathbf{n}_{\text {eff }}$ | PSUs |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | DEFF $_{\text {c }}$ | DEFF $_{\text {p }}$ | DEFF |  |  |  |  |  |
| Austria | Selection of individuals: Telephone book | 1.1 | 1.4 | 1.5 | 65 | 2,250 | 3,847 | 1,500 | 324 |
| Belgium | Selection of individuals: National register | 1.1 | 1.0 | 1.1 | 65 | 2,000 | 3,239 | 1,818 | 324 |
| Czech <br> Republic | Selection of households: SIPO database | 1.1 | 1.2 | 1.3 | 70 | 2,100 | 3,300 | 1,630 | 666 |
| Denmark | Selection of individuals: Central Person Register | 1.0 | 1.0 | 1.0 | 75 | 2,250 | 3,000 | 2,250 | - |
| Finland | Selection of individuals: Population register | 1.0 | 1.0 | 1.0 | 75 | 2,050 | 2,800 | 2,050 | - |
| France | Area based | 1.2 | 1.3 | 1.6 | 60 | 1,500 | 2,500 | 950 | 125 |
| Germany | Selection of individuals: Local residents registers | 1.4 | 1.1 | 1.5 | 70 | 3,066 | 4,868 | 2,017 | 163 |
| Greece | Area based | 1.1 | 1.2 | 1.3 | 70 | 2,100 | 3,100 | 1,570 | 438 |
| Hungary | Selection of individuals: Central register | 1.2 | 1.0 | 1.2 | 65 | 1,500 | 2,450 | 1,260 | 143 |
| Ireland | Selection of addresses: National Electoral Register | 1.2 | 1.3 | 1.6 | 63 | 2,400 | 4,233 | 1,500 | 220 |
| Israel | Selection of households: Telephone register | 1.2 | 1.3 | 1.6 | 70 | 2,394 | 3,600 | 1,506 | 200 |
| Italy | Selection of addresses: Electoral register | 1.1 | 1.0 | 1.1 | 70 | 2,000 | 3,000 | 1,820 | 125 |


|  | Frame | DEFF $_{\mathbf{c}}$ | DEFF $_{\mathbf{p}}$ | DEFF | Anticipated <br> Response <br> rate (\%) | $\mathbf{n}_{\text {net }}$ | $\mathbf{n}_{\text {gross }}$ | $\mathbf{n}_{\text {eff }}$ | PSUs |
| :--- | :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Luxembourg | Selection of households: Social security register | 1.0 | 1.40 | 1.40 | 30 | 1,510 | 5,033 | 1,080 | - |
| Netherlands | Selection of addresses: List of postal delivery points | 1.2 | 1.0 | 1.2 | 70 | 2,371 | 3,565 | 2,000 | - |
| Norway | Selection of individuals: National Population Register | 1.2 | 1.2 | 1.5 | 70 | 2,250 | 3,215 | 1,500 | 109 |
| Poland | Selection of individuals: Personal records of population | 1.1 | 1.0 | 1.1 | 71 | 2,000 | 2,978 | 1,718 | 200 |
| Portugal | Area based | 1.1 | 1.1 | 1.2 | 75 | 1,695 | 2,260 | 1,367 | 150 |
| Slovenia | Selection of individuals: Central register of population | 1.4 | 1.0 | 1.4 | 66 | 1,336 | 2,250 | 955 | 150 |
| Spain | Selection of individuals: Municipal rolls | 1.2 | 1.1 | 1.4 | 70 | 2,076 | 3,702 | 1,500 | 346 |
| Sweden | Selection of individuals: Population register | 1.0 | 1.0 | 1.0 | 75 | 2,198 | 3,000 | 2,198 | - |
| Switzerland | Selection of households: Telephone book | 1.2 | 1.2 | 1.5 | 40 | 2,205 | 6,660 | 1,500 | 220 |
| Turkey | Selection of households: Voters registries | 1.2 | 1.2 | 1.5 | 60 | 2,220 | 3,700 | 1,500 | 200 |
| UK | Selection of addresses: Postcode address files | 1.3 | 1.2 | 1.6 | 70 | 2,348 | 4,013 | 1,515 | 162 |

Appendix 2: Sampling issues in the "Specifications for participating countries"; Round 1 of the ESS

## Specification for the Survey

## 1.Population coverage

The survey will be representative of all persons aged 15 and over (no upper age limit) resident within private households in each country, regardless of their nationality, citizenship, language or legal status. [Please note that questionnaires are to be available in all languages spoken as a first language by 5 per cent or more of the population and interviewers must be available to administer them (see 5.12). For speakers of certain minority languages, however, it may be possible to use or adapt the questionnaire from another participating country.] Potential under-coverage of certain groups, say because of language problems or sampling frame deficiencies, or for any other reason, must be discussed with the CCT and their expert panel prior to deciding on the final sampling method, so that the problem can if at all possible be remedied.

## 2.The sample

The sample is to be selected by strict random probability methods at every stage and respondents are to be interviewed face-to-face (see section 5.13). Procedures for selecting a household from a multi-household address, and an individual within a household will be specified and agreed in advance. In any event, the relative selection probabilities of every sample member must be known and recorded, as should any remaining systematic noncoverage problems. Quota sampling is not permitted at any stage, nor is substitution of nonresponding households or individuals (whether 'refusals' or 'non-contacts').

## 3.Effective sample size

The minimum number of actual interviews to be achieved is $\mathbf{2 , 0 0 0}$ (except in countries whose total population is less then 2 million, when the minimum number is 1,000 ). Irrespective of the actual number of interviews, however, the minimum 'effective achieved sample size' should be 1,500, after discounting for design effects (see Appendix 1), or 800 in countries with populations of under 2 million. Thus, each country should determine the appropriate size of its initial issued sample by taking into account the realistic estimated impact of clustering, eligibility rates (where appropriate) and response rate on the effective sample size.

## 4.Over-sampling

Over-sampling (or using different selection probabilities) among certain subgroups or strata is acceptable provided that the total sample still complies with the effective sample size criterion. For instance, if a low response rate in certain strata (say large cities) is predicted, it would be efficient to over-sample in large cities by a factor because the resultant nonresponse weight would largely cancel out the selection probability weight, minimising any loss of precision due to weighting. Similarly, certain minority groups may be over-sampled to permit separate analysis of them. However, over-sampling will only help to ensure that certain cells are better filled than they otherwise would have been. It will not influence the target overall response rate. Moreover, since over-sampling will also necessitate subsequent additional weighting of the data to correct for the different selection probabilities, any proposals to over-sample must be discussed and agreed in advance with the CCT and their expert panel (see Appendix 1).

## 5.Documentation of sampling procedures

The precise sampling procedures to be employed in each country, and their implications for representativeness, must be documented in full and submitted in advance to the CCT for reference to the expert panel and 'signing off'. This precaution is to ensure that all countries within the ESS have defensible (and equivalent) national probability samples of their adult populations. The final sample design will be fully documented in the national technical report of the survey. It will include details of the definition and description of the sampling units used at each stage, the degree of clustering, any stratification factors applied to the sampling frame, any over-sampling, and an assessment of the ways in which selection probabilities might otherwise have varied at each stage. Such documentation will be translated into one or more variables within the national data file to indicate the relative selection probabilities of cases and to enable appropriate weighting strategies to be calculated.

## 6.Target response rates

Outcomes of all approaches to addresses, households and individuals in the sample will be defined and recorded according to a pre-specified set of categories that distinguish noneligibility, non-contacts and refusals. The proportion of non-contacts should not exceed 3 per cent of all sampled units, and the minimum target response rate - after discounting ineligibles (and other 'deadwood', as defined by the CCT) - should be 70\%. This figure is likely to be exceeded in certain countries and the ESS as a whole would be damaged if major national variations in response rates were to occur. Survey organisations should thus cost their surveys with this response rate in mind and consider what steps may be required to achieve it.

## Appendix 3: Appendix 1 of the "Specifications for participating countries" (Rules for estimating design effects)

## Effective Sample Size

The effective sample size (neff) is the size of a simple random sample which would produce the same precision (standard errors) as the design actually used. Typically, neff is less than the actual number of achieved interviews, m, as certain aspects of survey design - for example, clustering or the use of differing selection probabilities - tend to reduce the precision of estimates. The reduction of precision is known as the design effect (DEFF):
$D E F F$ = Actual sampling variance / Sampling variance with simple random samples of same size;
$D E F F=m /$ neff, so neff $=m / D E F F$
We therefore need to be able to predict the value of DEFF for a proposed sample design, in order to determine how many interviews should be achieved so as to produce a particular value of neff. We suggest that two components of $D E F F$ should be taken into account at the design stage - the design effect arising from differing selection probabilities $\left(D E F F_{p}\right)$ and the design effect arising from clustering $\left(D E F F_{c}\right)$. Then $D E F F=D E F F_{p} \times D E F F_{c}$. We then also need to predict the survey response rate (and the proportion of ineligibles on the sampling frame, if relevant) in order to determine the size of the initial sample ( n ) required in order to achieve approximately m interviews.

## Design Effects due to Differing Selection Probabilities

In some countries which have accessible population registers, it will be possible to select an equal-probability sample from the survey population. In other countries, it will be necessary to select the sample in stages, with the penultimate stage being residential addresses. In this case, each person's selection probability will depend on their household size. Another reason why differing selection probabilities might be used is if important minority groups were to be over-sampled.

If differing selection probabilities are to be used - for whatever reason - the associated design effect should be predicted. This can be done very simply, using the following formula

$$
D E F F_{p}=\frac{m\left(\sum_{i} m_{i} w_{i}^{2}\right)}{\left(\sum_{i} m_{i} w_{i}\right)^{2}}
$$

where there are $\mathrm{m}_{\mathrm{i}}$ respondents in the $\mathrm{i}^{\text {th }}$ selection probability class, each receiving a weight of $\mathrm{w}_{\mathrm{i}}$, where $\alpha$ means 'proportional to'

$$
w_{i} \alpha \frac{N_{i}}{m_{i}}
$$

(This formula assumes that the population variance of survey variables will not vary over selection probability classes - a reasonable assumption in most situations)

It is anticipated that in most countries it will be efficient to select a multi-stage, clustered, sample. In such situations there will also be a design effect due to clustering:
$D E F F_{c}=1+(b-1) \rho$
where $b$ is the mean number of respondents per cluster and $\rho$ is the intra-cluster correlation (or "rate of homogeneity") - a measure of the extent to which persons within a clustering unit are more homogeneous than persons within the population as a whole (see Kish, 1994, Survey Sampling, pp. 161-164 (New York: Wiley and Sons, Inc.)). This design effect can be estimated, at least crudely, from knowledge of other surveys and/or the nature of the clustering units.

In practice, all elements of the overall design effect, including that due to differing selection probabilities and that due to clustering, will take different values for different survey estimates. For sample design purposes, an average value should be used.

## Example: How to determine the size of issued sample

We have prescribed neff $>1500$.
To determine m , we must first estimate $\quad D E F F=D E F F_{p} \times D E F F_{c}$

1. Suppose the proposed clustering units are administrative areas of around 5,000 households on average and that based on data from other surveys, we expect that for these areas, $\rho$ will take values of around 0.02 for many variables. Then, if we are proposing a design with a mean of 15 interviews per cluster:
$D E F F_{c}=1+(15-1) \times 0.02=1.28$.
[Note: If there is no available empirical evidence at all upon which to base an estimate of $\Delta$, then we suggest that a value of 0.02 should be used.]
2. Suppose that the only available sampling frame is a list of addresses and that these must be selected with equal probabilities. The proposed design is then randomly to select one person to interview at each address. This is the only aspect of the proposed design that involves differing selection probabilities. Then, we can use population statistics on the distribution of household size to estimate the number of respondents in each selection probability class, thus:


|  |  |  | $\mathbf{w}_{\mathbf{i}}$ | $\mathbf{m}_{\mathbf{i} \mathbf{w}} \mathbf{i}$ | $\mathbf{m}_{\mathbf{i}} \mathbf{w}_{\mathbf{i}}{ }^{\mathbf{2}}$ |
| :--- | :--- | :--- | :--- | :--- | :--- |
| 1 | 0.35 | 0.35 m | 1 | 0.35 m | 0.35 m |
| 2 | 0.45 | 0.45 m | 2 | 0.90 m | 1.80 m |
| 3 | 0.12 | 0.12 m | 3 | 0.36 m | 1.08 m |
| 4 | 0.06 | $0,06 \mathrm{~m}$ | 4 | 0.24 m | 0.96 m |
| 5 | 0.02 | 0.02 m | 5 | 0.10 m | 0.50 m |

The population distribution of household size appears in the first two columns. From this, we can predict that the sample distribution will be as shown in the third column. We can thus predict $\mathrm{DEFF}_{\mathrm{p}}$ :

$$
D E F F_{p}=m \times 4.69 m /(1.95 m)^{2} \quad=\quad 4.69 / 1.95^{2}=1.23
$$

3. Thus, we predict DEFF $=1.28 \times 1.23=1.57$. Consequently, to achieve neff $>1,500$ with this design, we would need $\mathrm{m}>1,500 \times 1.57=2,355$.
4. The final stage is to calculate the sample size to select initially in order to be likely to achieve around 2,355 interviews. Suppose we anticipate a response rate of $80 \%$ and that $5 \%$ of the sampling frame units will be ineligible (e.g. addresses which do not contain a resident household), then:
$\mathrm{n}=(\mathrm{m} / 0.80) / 0.95=3,098$
So we would select a sample of at least 3,100 addresses.

[^0]:    ${ }^{1}$ see www.europeansocialsurvey.org (Technical annex).
    2 "The precise sampling procedures to be employed in each country, and their implications for representativeness, must be documented in full and submitted in advance to the CCT for reference to the expert panel and 'signing off'." see www.europeansocialsurvey.org, Specifications for Participating countries.

[^1]:    ${ }^{3}$ see www.europeansocialsurvey.org
    ${ }^{4}$ In countries in which any minority language is spoken as a first language by $5 \%$ or more of the population, the questionnaire will be translated into that language.

[^2]:    ${ }^{5}$ For face to face studies frequently clustered designs are used because they are usually less expansive (shorter ways for the interviewers) then unclustered designs.
    ${ }^{6}$ Of course, there are also other influences on the data quality besides those because of sampling - such as features of the interviewers, question wording, translation problems and so on. However, we cannot deal with them in this paper.
    ${ }^{7}$ Susan Purdon had to leave the panel in August, 2002. From that point in time the panel consisted of four experts.

[^3]:    Special Features of the design

    The need to estimate selection probabilities from the interview data, as described above.

[^4]:    Special Features of the No over-sampling of specific groups. design

[^5]:    Special Features of the design

    No over-sampling in Great Britain.
    Funding has just been confirmed for a boosted sample in Northern Ireland, the aim being to achieve an additional 950 cases there.

